

**An Application
of the Closed Testing Principle
to Enhance One-Sided Confidence Regions
for a Multivariate Location Parameter**

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Outline

Introduction

Direct Derivation of Confidence Regions for ϑ

Enhanced Confidence Regions for ϑ

One-Sided Confidence Regions

- $\mathbf{X}_1, \dots, \mathbf{X}_n$ i. i. d. random vectors in \mathbb{R}^p
- $\mathbf{X}_i \sim P_{\vartheta}$, $\vartheta \in \Theta$ unknown
- P_{ϑ} (at least directionally) symmetric w. r. t. ϑ

Problem: Find a $1 - \alpha$ confidence region for ϑ that is

- as strict as possible in specific directions
- possibly unbounded in “irrelevant” directions

(e. g. a cone or an orthant).

Connection with One-Sided Location Tests

Let φ_α be a non-randomized level α test for

$$H_0 : \vartheta \in \Theta_0(\gamma) \quad \text{vs.}$$

$$H_1 : \vartheta \in \Theta \setminus \Theta_0(\gamma).$$

(E. g. $\Theta_0(\gamma) = \gamma + (-\infty, 0]^p$)

Inversion of φ_α

$\Rightarrow \mathcal{C}_{1-\alpha}(\mathbf{X}_1, \dots, \mathbf{X}_n) = \{\gamma : \varphi_\alpha((\mathbf{X}_1, \dots, \mathbf{X}_n), \gamma) = 0\}$, and

$$P_{\vartheta}(\mathcal{C}_{1-\alpha}(X) \ni \gamma) \geq 1 - \alpha \quad \forall \vartheta \in \Theta_0(\gamma) \quad \forall \gamma \in \Theta.$$

$\mathcal{C}_{1-\alpha}$ is a $1 - \alpha$ confidence region for the meta-parameter γ .

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Then

$$P_{\vartheta}(\mathcal{C}_{1-\alpha}(X) \ni \vartheta) \geq 1 - \alpha \quad \forall \vartheta \in \Theta.$$

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Problems:

- conservative
- unpleasant shape – to be illustrated...

Min and Max Tests

Min Test

Reject $H_0 : \exists j \in \{1, \dots, p\} : \vartheta_j \leq \gamma_j$ in favor of $H_1 : \vartheta > \gamma$ at the level α if and only if

$$\varphi_{j,\alpha}((X_{1j}, \dots, X_{nj}), \gamma_j) = 1, \forall j \in 1, \dots, p.$$

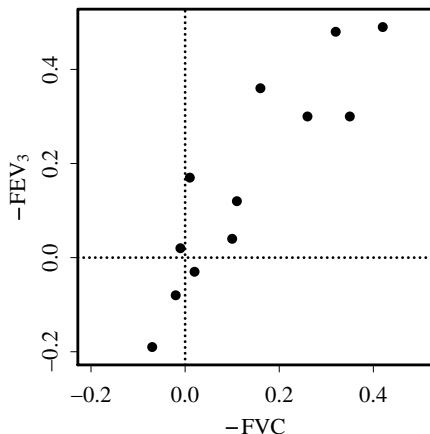
Bonferroni Max Test

Reject $H_0 : \vartheta \leq \gamma$ in favor of $H_1 : \exists j \in \{1, \dots, p\} : \vartheta_j > \gamma_j$ at the level α if and only if

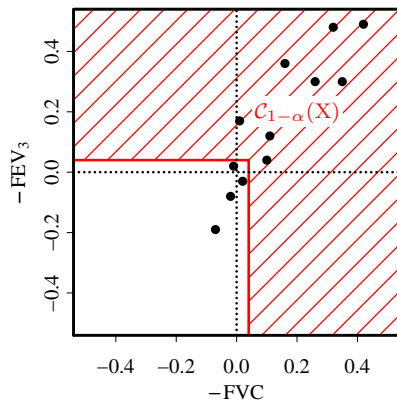
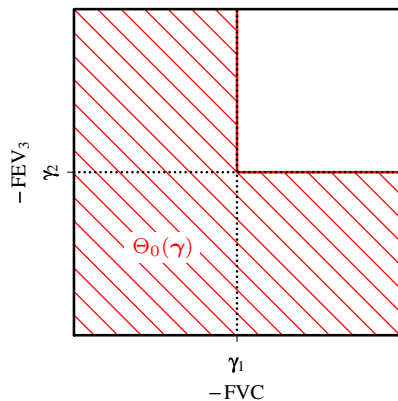
$$\exists j \in 1, \dots, p : \varphi_{j,\alpha/p}((X_{1j}, \dots, X_{nj}), \gamma_j) = 1.$$

Example

Two variables of the pulmonary function data by Randles (1989) (slightly modified from Merchant et al., 1975).

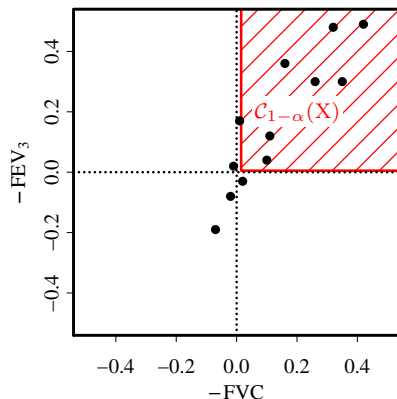
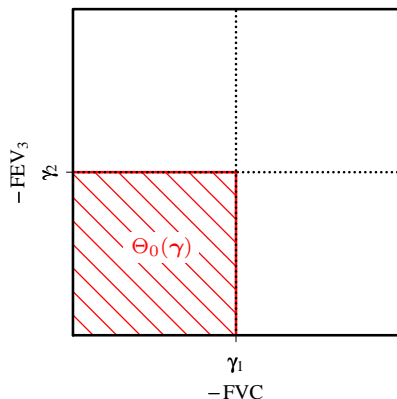


Pulmonary Function Data, Wilcoxon Min Test



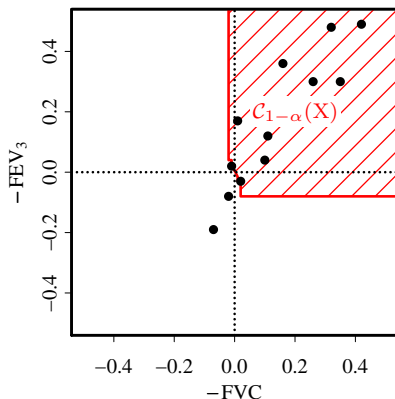
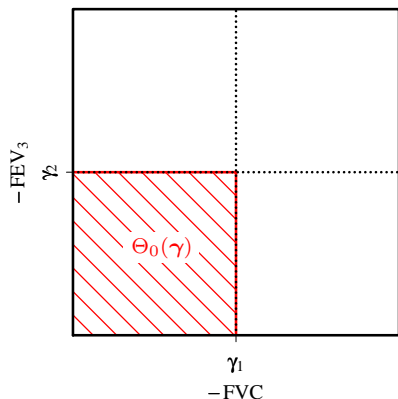
$$C_{1-\alpha}(X) = c_{1-\alpha}(X) - \Theta_0(\mathbf{0})$$

P. F. Data, Wilcoxon Bonferroni Max Test



$$C_{1-\alpha}(X) = c_{1-\alpha}(X) - \Theta_0(\mathbf{0})$$

P. F. Data, Sign Test by Larocque/Labarre (2004)



$$\mathcal{C}_{1-\alpha}(X) \approx \mathbf{c}_{1-\alpha}(X) - \Theta_0(\mathbf{0}) \text{ (outside a sufficiently large ball)}$$

Back to the Drawbacks of the Direct Approach

If confidence regions for γ are directly used as confidence regions for ϑ , they are

- usually conservative and
- similar in shape to $-\Theta_0$, rather than to $\Theta_1 = \Theta \setminus \Theta_0$.

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Enhanced Confidence Regions for ϑ (1)

$$\text{Temptation: } \tilde{\mathcal{C}}_{1-\alpha}(\mathbf{X}) = \bigcap_{\gamma \notin \mathcal{C}_{1-\alpha}(\mathbf{X})} \Theta_1(\gamma)$$

Enhanced Confidence Regions for ϑ (1)

Temptation: $\tilde{\mathcal{C}}_{1-\alpha}(\mathbf{X}) = \bigcap_{\gamma \notin \mathcal{C}_{1-\alpha}(\mathbf{X})} \Theta_1(\gamma)$

\Rightarrow liberal – multiple testing problem!

Solution: Reduce the set of possible meta-parameters in advance.

Enhanced Confidence Regions for ϑ (2)

Let $\mathcal{C}_{1-\alpha} : \mathcal{X} \rightarrow \mathcal{P}(\mathbb{R}^p)$ be a $1 - \alpha$ confidence region for γ based on $(\Theta_0(\gamma))_{\gamma \in \mathbb{R}^p}$.

Let $\Theta_0(\gamma) = \gamma + \Theta_0(\mathbf{0})$, $\forall \gamma \in \mathbb{R}^p$, closed, $\Theta_1(\gamma) = \mathbb{R}^p \setminus \Theta_0(\gamma)$.

Assume that $\Theta_0(\gamma) \subset \Theta_0(\gamma + (\delta, \dots, \delta)^T)$, $\forall \gamma \in \mathbb{R}^p$, $\delta > 0$.

With $\gamma_i = (i, \dots, i)^T \in \mathbb{R}^p$, $\forall i \in I = [\ell, \infty)$, define

$$\tilde{\mathcal{C}}_{1-\alpha}(X) := \bigcap_{i \in I: \gamma_{i'} \notin \mathcal{C}_{1-\alpha}(X) \forall i' \leq i} \Theta_1(\gamma_i).$$

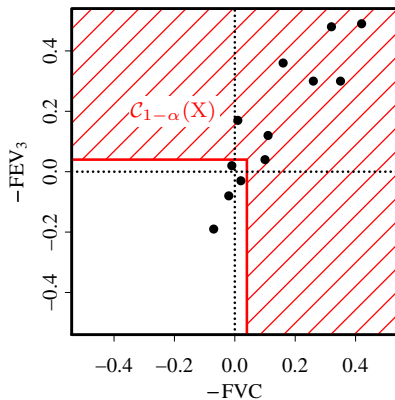
Then

$$P_{\vartheta} \left(\tilde{\mathcal{C}}_{1-\alpha}(X) \ni \vartheta \right) \geq 1 - \alpha \quad \forall \vartheta \in \mathbb{R}^p.$$

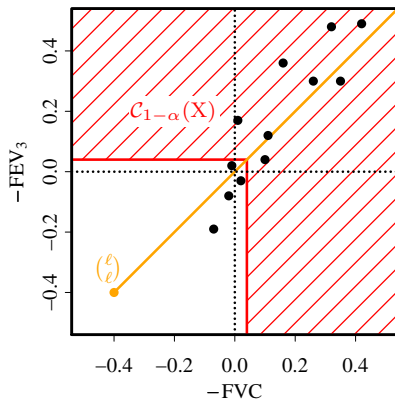
Idea of the Proof

- $(\Theta_0(\gamma_i))_{i \in I}$ is closed under (finite and infinite) intersections.
- Apply the closed testing principle (Marcus, Peritz, and Gabriel, 1976).
- Translate to confidence regions.

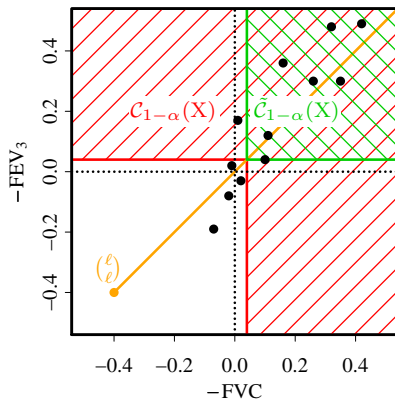
Pulmonary Function Data, Wilcoxon Min Test



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Pulmonary Function Data, Wilcoxon Min Test



Enhanced Confidence Regions: Properties

- $\tilde{\mathcal{C}}_{1-\alpha}(\mathbf{X}) = \mathbf{c}_{1-\alpha}(\mathbf{X}) + \Theta_1(\mathbf{0})$ (by definition)
- $\tilde{\mathcal{C}}_{1-\alpha}(\mathbf{X}) \subset \mathcal{C}_{1-\alpha}(\mathbf{X})$ under suitable conditions ($\Theta_1(\mathbf{0})$ convex cone, translation invariance, and a monotonicity property)

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Disadvantage: Restricted set of possible confidence regions (search on the diagonal)

Summary

- A confidence region $\mathcal{C}_{1-\alpha}(\mathbf{X})$ obtained by inversion of a test for a composite null hypothesis is for a meta-parameter.
- Even if $\mathcal{C}_{1-\alpha}(\mathbf{X})$ may also be a confidence region for the parameter ϑ itself, it is not very useful.
- The proposed method based on the closed testing principle yields a confidence region $\tilde{\mathcal{C}}_{1-\alpha}(\mathbf{X})$ with a more useful shape.
- $\tilde{\mathcal{C}}_{1-\alpha}(\mathbf{X})$ is also less conservative than $\mathcal{C}_{1-\alpha}(\mathbf{X})$ under suitable conditions.

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