

# Multiple Testing in Change-Point Problem with Application to Safety Signal Detection

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Introduction

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Two-Sequence Change-Point  
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# Overview

## Introduction

Concept/applications

Classification

Methods

## Change-Point Analysis

One changepoint

Multiple Changepoints

## Multiple Testing

Šidák inequality

Closure principle

Proposed method

## Two-Sequence Change-Point Problem

## Example

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# Introduction (1)

## Introduction

- **Concept/applications**
- Classification
- Methods

## Change-Point Analysis

## Multiple Testing

## Two-Sequence Change-Point Problem

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## Summary

- Change-point analysis concerns with the inference on the point(s) in a sequence of random process at which the distribution changes
- Applications
  - Statistical quality control – Online detection of changes in quality operations
  - Public Health – Sequentially monitoring the number of cases of a disease for potential outbreak
  - Medicine – Post-marketing surveillance, dose-finding
  - Biomedical signal processing – Online detection of biomedical signals such as Electroencephalogram (EEG) and electrocardiogram (ECG)
  - Meteorology – Global warming
  - Finance – Detection of business cycles
  - Seismology



## Introduction (2)

### Introduction

- Concept/applications
- **Classification**
- Methods

### Change-Point Analysis

### Multiple Testing

### Two-Sequence Change-Point Problem

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## Classifications of change-point analysis

- Continuous versus discrete
- Retrospective (fixed sample) versus prospective (online sequential)
- Parametric versus non-parametric
- Frequentist versus Bayesian
- One change-point versus multiple change-points



# Introduction (3)

## Introduction

- Concept/applications
- Classification
- **Methods**

## Change-Point Analysis

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## Statistical methods for change-point analysis

- Likelihood ratio procedures for parametric models
- Non-parametric methods – Mann-Whitney  $U$  test, Wilcoxon rank test
- Regression-based methods (including curve fitting)
- Cumulative sum (CUSUM) methods
- Bayesian analysis and its variations
- Sequential methods
- information criterion
- Wavelet transformation



# One Change-Point Problem (1)

Introduction

Change-Point Analysis

● One changepoint

● Multiple Changepoints

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Summary

- Let  $(X_1, \dots, X_T)$  be a sequence of independent random variables, ordered in time interval, each with density function  $f(X_i|\mu_i)$  where  $E(X_i) = \mu_i, i = 1, \dots, T$
- Consider the model for one change-point in means at time interval  $\tau$
- The null hypothesis

$$H_0 : \mu_1 = \dots = \mu_T$$

against against

$$H_1 : \mu_1 = \dots = \mu_\tau \neq \mu_{\tau+1} = \dots = \mu_T$$

for an unknown  $\tau$



## One Change-Point Problem (2)

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Summary

- Let  $\mu_1 = \dots = \mu_\tau = \mu'_0$  with known  $\mu'_0$  before the change-point  $\tau$  and  $\mu_{\tau+1} = \dots = \mu_T = \mu'_1$  with known  $\mu'_1$  after the change-point. Then the log likelihood function is

$$\ell(\tau) = \sum_{i=1}^{\tau} \log f(X_i | \mu'_0) + \sum_{i=\tau+1}^T \log f(X_i | \mu'_1) \quad (1)$$

- The log likelihood ratio test statistic for testing  $H_0$  against  $H_1$  is

$$\log \ell(\tau) = \sum_{i=\tau+1}^T \log f(X_i | \mu'_0) - \sum_{i=\tau+1}^T \log f(X_i | \mu'_1) \quad (2)$$



## One Change-Point Problem (3)

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● One changepoint

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- Traditionally, the null hypothesis  $H_0$  of no change-point against  $H_1$  of one change-point over the  $T$  time intervals is rejected if

$$2 \sup_{\tau} \log \ell(\tau) > \chi_{\alpha,1}^2 \quad (3)$$

- The MLE  $\hat{\tau}$  of  $\tau$  is obtained by maximizing (1)
- When  $\mu'_0$  and  $\mu'_1$  are unknown, the MLE's  $\hat{\mu}'_0$ ,  $\hat{\mu}'_1$  and  $\hat{\tau}$  can be obtained by simultaneously maximizing (1) w.r.t.  $\mu'_0$ ,  $\mu'_1$  and  $\tau$





# Multiple Change-Points (1)

Introduction

Change-Point Analysis

- One changepoint
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- The test continues sequentially for testing  $H_{m-1}$  of  $m - 1$  change-points against  $H_m$  of  $m$  change-points,  $m = 1, \dots, T - 1$ , until an acceptance occurs
- That is, the test starts from  $m = 0$  (against  $m = 1$ ) towards  $m = T - 1$

$$H_0 : \mu_1 = \dots = \mu_T$$

$$H_1 : \mu_1 = \dots = \mu_\tau \neq \mu_{\tau+1} = \dots = \mu_T$$

$$H_2 : \mu_1 = \dots = \mu_{\tau_1} \neq \mu_{\tau_1+1} = \dots = \mu_{\tau_2} \neq \mu_{\tau_2+1} = \dots = \mu_T$$

$$\vdots \quad \vdots$$

$$H_{T-1} : \mu_1 \neq \dots \neq \mu_T$$



## Multiple Change-Points (2)

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● One changepoint

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Summary

- Binary segmentation (Vostrikova 1981)
  - Test for no change point  $H_0$  against one change-point  $H_1$
  - If  $H_0$  is rejected, test the two subsequences before and after the change-point identified in the above step separately for a change
  - Repeat the process until no change-points are found in any of the subsequences
  - The collection of change-points identified from the above steps are  $\{\hat{\tau}_1, \dots, \hat{\tau}_k\}$  and the estimated number of change-points is then  $k$



## Multiple Change-Points (3)

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Change-Point Analysis

- One changepoint
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A hierarchy of sub-hypothesis tests (Hogg 1961)

- Let  $\Omega$  denote the total parameter space and  $\Omega^*$  a subspace of  $\Omega$ .
- It is desired to test  $H'_0 : \theta \in \Omega^*$  against  $H'_1 : \theta \in \Omega - \Omega^*$ .
- Suppose there are certain intermediate hypotheses. Let  $\Omega_i$  be a subset of  $\Omega_{i-1}$ ,  $i = 1, \dots, t - 1$ , such that

$$\Omega = \Omega_0 \supset \Omega_1 \supset \dots \supset \Omega_{t-1} = \Omega^*$$

where each  $\Omega_i$  corresponds to an intermediate hypothesis

- Testing  $H'_0$  against  $H'_1$  can be carried out by iteratively testing the following hypotheses:

$$H_0^i : \theta \in \Omega_i \text{ versus } H_1^i : \theta \in \Omega_{i-1} - \Omega_i,$$

$$i = 1, \dots, t - 1$$



# Multiple Change-Points (4)

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- One changepoint
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A hierarchy of sub-hypothesis tests (cont'd)

- To test  $H'_0$  against  $H'_1$ , we first test

$$H_0^1 : \theta \in \Omega_1 \text{ against } H_1^1 : \theta \in \Omega_0 - \Omega_1$$

- If  $H_0^1$  is accepted, we then test

$$H_0^2 : \theta \in \Omega_2 \text{ against } H_1^2 : \theta \in \Omega_1 - \Omega_2$$

- In general, if  $H_0^{i-1}$  is accepted, we continue to test

$$H_0^i : \theta \in \Omega_i \text{ against } H_1^i : \theta \in \Omega_{i-1} - \Omega_i$$

- $H'_0$  is rejected if any one of  $H_0^1, \dots, H_0^{t-1}$  is rejected
- $H'_0$  is accepted if and only if all of  $H_0^1, \dots, H_0^{t-1}$  are accepted



# Multiple Testing in Change-Point Analysis (1)

Introduction

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Multiple Testing

• Šidák inequality

• Closure principle

• Proposed method

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Summary

- “If you torture the data long enough, it will confess anything you want” – Nobel Laureate Ronald Coase
- Generalized likelihood ratio (GLR) test
  - Let  $\lambda_i = L(\hat{\Omega}_i) / L(\hat{\Omega}_{i-1})$  be the likelihood ratio for testing  $H_0^i$  against  $H_1^i$ ,  $i = 1, \dots, t - 1$
  - The GLR for  $H_0'$  against  $H_1'$  is given by

$$\lambda = \frac{L(\hat{\Omega}_{t-1})}{L(\hat{\Omega}_0)} = \prod_{i=1}^{t-1} \frac{L(\hat{\Omega}_i)}{L(\hat{\Omega}_{i-1})} = \prod_{i=1}^{t-1} \lambda_i \quad (4)$$

- The  $\lambda_i$ 's are mutually stochastically independent test statistics
- Significance level for each test  $\alpha_i = 1 - (1 - \alpha)^{1/(t-1)}$ , where  $\alpha$  is the family-wise type I error rate



# Multiple Testing in Change-Point Analysis (2)

Introduction

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- Šidák inequality
- **Closure principle**
- Proposed method

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## Closure principle

- Suppose  $T = 4$ , then there exist at most  $m = 3$  change-points.

Reformulation of the  $H_i$ 's

$$H_0 : \{ \mu_1 = \mu_2 = \mu_3 = \mu_4 \}$$

$$H_1 : \{ \mu_1 = \mu_2 = \mu_3 \}, \{ \mu_2 = \mu_3 = \mu_4 \}, \\ \{ \mu_1 = \mu_2 \} \cap \{ \mu_3 = \mu_4 \}$$

$$H_2 : \{ \mu_1 = \mu_2 \}, \{ \mu_2 = \mu_3 \}, \{ \mu_3 = \mu_4 \}$$

$$H_3 : \{ \mu_1 \}, \{ \mu_2 \}, \{ \mu_3 \}, \{ \mu_4 \}$$

- This forms the closure of the family by taking all possible intersections
- The closed family resembles the one that is formed for all pair-wise comparisons, but much smaller; it consists of hypotheses of homogeneity of successive means and their intersections



## Multiple Testing in Change-Point Analysis (3)

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- Šidák inequality
- Closure principle
- Proposed method

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Closed test procedure for dose-response (Marcus, et al, 1976; Begun and Gabriel, 1981; Rom, et al, 1994) – Let  $k$  denote the cardinality of a subset homogeneity hypothesis  $H_K$

- Reject a subset homogeneity hypothesis  $H_K$  at level  $\alpha_k = \alpha k/t$
- Retain  $H_K$  at level  $\alpha$
- Otherwise, if  $H_K$  is rejected at level  $\alpha$  but not at level  $\alpha_k$ , then  $H_K$  is rejected if every hypothesis  $H_R$  that concerns means in the complement of  $K$  is rejected at level  $\alpha_r$



# Multiple Testing in Change-Point Analysis (4)

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- Closure principle
- Proposed method

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Proposed modification of the above closed test procedure

- Reject  $H_0^i$ 
  - If  $\Omega_i$  is not contained in any accepted set, and
  - If  $H_0^i$  is rejected at level  $\alpha_i = 1 - (1 - \alpha)^{(t-1-i)/(t-1)}$
- Retain  $H_0^i$ 
  - If  $\Omega_i$  is contained in another accepted set, or
  - If  $H_0^i$  is not rejected at level  $\alpha_i = 1 - (1 - \alpha)^{(t-1-i)/(t-1)}$





# Two-Sequence Change-Point Problem

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- Suppose that there are two independent sequences of random processes  $X_{i1}$  and  $X_{i2}$ , with  $X_{ij} \sim f(\mu_{ij})$ ,  $i = 1, \dots, T$  and  $j = 1, 2$
- The question of interest is whether there are an abrupt change in the ratios of the two random variables across the time period
- Let  $\gamma_i = \mu_{i1}/\mu_{i2}$ ,  $i = 1, \dots, T$ . Then this is equivalent to simultaneously testing the null hypotheses

$$H_0 : \gamma_1 = \dots = \gamma_T$$

against  $H_1$  : there is at least one change point



# An Example (1)

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● Example

Summary

## Pharmacovigilance and Post-Marketing Safety Surveillance

- An integrated part of biopharmaceutical development
- Activities involve in detection, assessment, understanding and prevention of adverse effects and any drug-related problems
- Pharmacovigilance plan (PvP)
  - As part of Marketing Authorization Application, the PvP must be prepared in compliance with regulatory request on potential safety impact of product modification
  - The PvP describes routine pharmacovigilance practice, as well as special action plan including Post-marketing Safety Surveillance Analysis (PSSA)
  - The PSSA will evaluate all potential safety signals, with special attention to proportional change of a particular adverse event (system) over time



## An Example (2)

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● Example

Summary

### PSSA

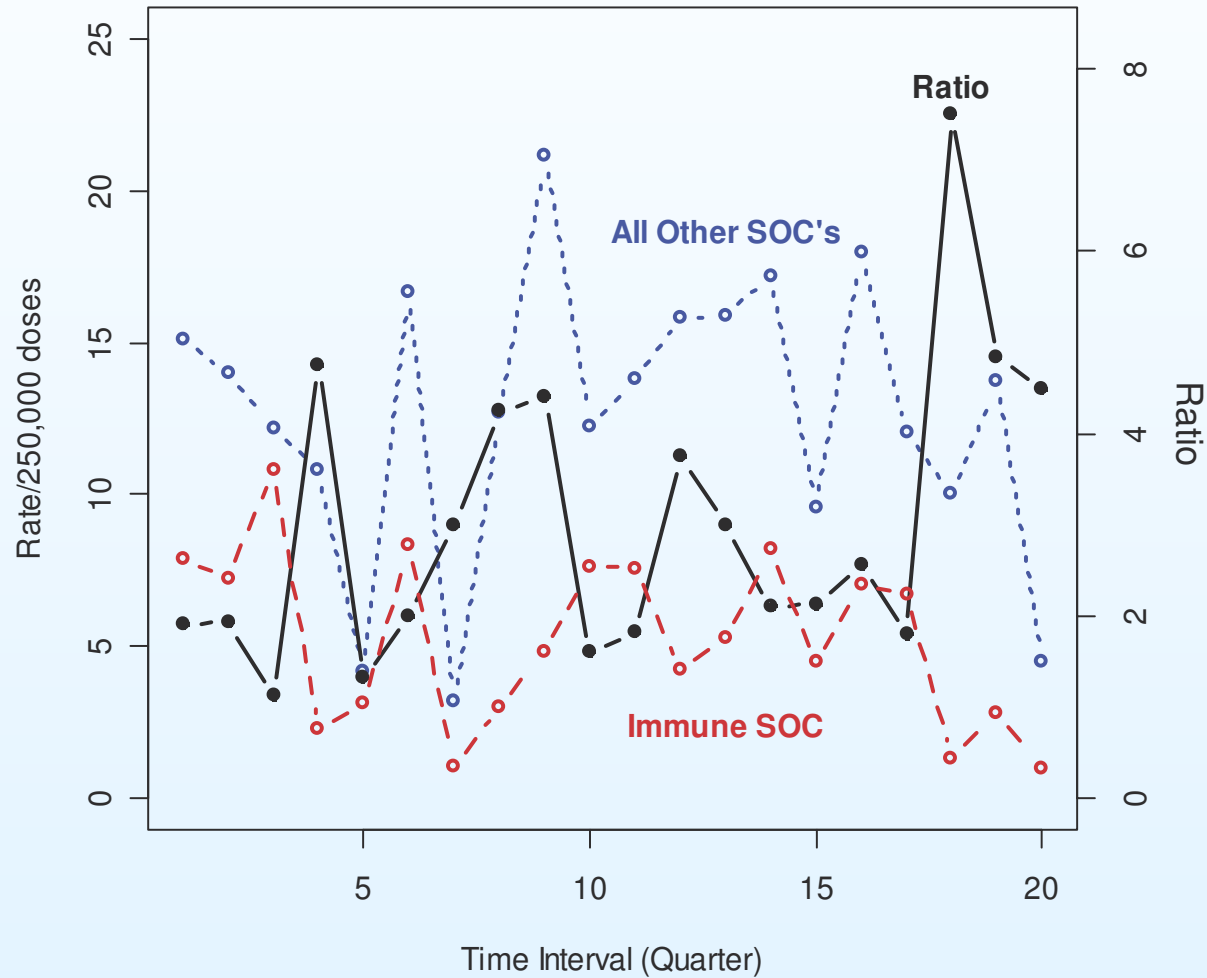
- Uses spontaneous reporting databases (AERS, VAERS, company's spontaneous reporting databases, etc.) and other epidemiological studies
- Data mining tools – Empirical Bayes, neural network, etc.
- Proportional change over time – useful to detect the impact of drug modification on the reporting of a particular (body system) adverse experience. For example,
  - Name change
  - Combination of two or more independent vaccines
  - New technology  $\Rightarrow$  manufacturing process change
  - ...



# An Example (3)

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- Example
  - Example
- Summary

Figure 1: AE reporting rates for autoimmune SOC and other SOC's



## An Example (4)

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● Example

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Estimates of the number of change-points:

- All computations are carried out using compiled R functions and SAS<sup>®</sup> macros
- No multiplicity adjustment
  - Each tested at  $\alpha = 0.05$
  - 4 change-points: 3, 4, 9, and 17
- Šidák inequality
  - Each tested at  $\alpha^* = 1 - (1 - 0.05)^{1/19} = 0.0027$
  - 1 change-point: 17
- Proposed method
  - Each  $\lambda_i$  is tested at  $0.05 \times (19 - i)/19$
  - 2 change-points: 3 and 17



# Summary

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Summary

- Exact method for computing rejection probabilities
- Simulations
- Constant rate within interval
- multivariate change-point problem
- Sequential or online change-point problem
- Continuous-time estimate of change-point
- Bayesian attempt
- Potential biases in post-marketing spontaneous data reporting

